### 5.2.4.1 Least Squares Approach

For illustrative purposes, suppose $N=2$ and $\mathrm{J}=4$ so that

$$
\begin{array}{lll}
y_{11}=\theta_{1} \phi_{1}+\varepsilon_{11} & ; & y_{21}=\theta_{2} \phi_{1}+\varepsilon_{21} \\
y_{12}=\theta_{1} \phi_{2}+\varepsilon_{12} & ; & y_{22}=\theta_{2} \phi_{2}+\varepsilon_{22} \\
y_{13}=\theta_{1} \phi_{3}+\varepsilon_{13} & ; & y_{23}=\theta_{2} \phi_{3}+\varepsilon_{23} \\
y_{14}=\theta_{1} \phi_{4}+\varepsilon_{14} & ; & y_{24}=\theta_{2} \phi_{4}+\varepsilon_{24}
\end{array}
$$

Step (a): Treating $\phi_{j}$ as known, we can re-write model (1) in matrix form

$$
\begin{equation*}
\mathbf{y}=\mathbf{x}^{T} \boldsymbol{\theta}+\boldsymbol{\varepsilon} \tag{2}
\end{equation*}
$$

Where $\quad \mathbf{y}=\left(\begin{array}{l}y_{11} \\ y_{12} \\ y_{13} \\ y_{14} \\ y_{21} \\ y_{22} \\ y_{23} \\ y_{24}\end{array}\right), \quad \mathbf{x}=\left(\begin{array}{cc}\phi_{1} & 0 \\ \phi_{2} & 0 \\ \phi_{3} & 0 \\ \phi_{4} & 0 \\ 0 & \phi_{1} \\ 0 & \phi_{2} \\ 0 & \phi_{3} \\ 0 & \phi_{4}\end{array}\right), \quad$ and $\boldsymbol{\theta}=\binom{\theta_{1}}{\theta_{2}}$
$\boldsymbol{\varepsilon}$ is a vector of random errors

The least squares system of normal equations yields

$$
\hat{\boldsymbol{\theta}}=\left(\mathbf{x}^{T} \mathbf{x}\right)^{-1} \mathbf{x}^{T} \mathbf{y}
$$

$$
\begin{aligned}
& \mathbf{x}^{T} \mathbf{x}=\left[\begin{array}{cccccc}
\phi_{1} \phi_{2} \phi_{3} \phi_{4} & 0 & 0 & 0 & 0 \\
0 & 0 & 0 & 0 & \phi_{1} \phi_{2} & \phi_{3} \phi_{4}
\end{array}\right]\left[\begin{array}{ll}
\phi_{1} & 0 \\
\phi_{2} & 0 \\
\phi_{3} & 0 \\
\phi_{4} & 0 \\
0 & \phi_{1} \\
0 & \phi_{2} \\
0 & \phi_{3} \\
0 & \phi_{4}
\end{array}\right]=\left[\begin{array}{ll}
\sum_{j=1}^{4} \phi_{j}^{2} & 0 \\
0 & \sum_{j=1}^{4} \phi_{j}^{2}
\end{array}\right] \\
& \mathbf{x}^{T} \mathbf{y}=\left[\begin{array}{l}
\phi_{1} y_{11}+\phi_{2} y_{12}+\phi_{3} y_{13}+\phi_{4} y_{14} \\
\phi_{1} y_{21}+\phi_{2} y_{22}+\phi_{3} y_{23}+\phi_{4} y_{24}
\end{array}\right]=\left[\begin{array}{l}
\sum_{j=1}^{4} \phi_{j} y_{1 j} \\
\sum_{j=1}^{4} \phi_{j} y_{2 j}
\end{array}\right] \\
& \text { So } \hat{\boldsymbol{\theta}}=\left(\mathbf{x}^{T} \mathbf{x}\right)^{-1} \mathbf{x}^{T} \mathbf{y}=\frac{1}{\sum_{j=1}^{4} \phi_{j}^{2}}\left[\begin{array}{l}
\sum_{j=1}^{4} \phi_{j} y_{1 j} \\
\sum_{j=1}^{4} \phi_{j} y_{2 j}
\end{array}\right] \\
& \operatorname{Var}(\hat{\boldsymbol{\theta}})=\sigma^{2}\left(\mathbf{x}^{T} \mathbf{x}\right)^{-1}=\frac{\sigma^{2}}{\sum_{j=1}^{4} \phi_{j}^{2}}\left[\begin{array}{ll}
1 & 0 \\
0 & 1
\end{array}\right]
\end{aligned}
$$

Step (b): Now treating $\boldsymbol{\theta}_{i}$ as known, we can re-write model (1) in matrix form

$$
\begin{equation*}
\mathbf{y}=\mathbf{x} \varphi+\varepsilon \tag{3}
\end{equation*}
$$

Where $\mathbf{y}=\left[\begin{array}{l}y_{11} \\ y_{12} \\ y_{13} \\ y_{14} \\ y_{21} \\ y_{22} \\ y_{23} \\ y_{24}\end{array}\right], \quad \mathbf{x}=\left[\begin{array}{cccc}\theta_{1} & 0 & 0 & 0 \\ 0 & \theta_{1} & 0 & 0 \\ 0 & 0 & \theta_{1} & 0 \\ 0 & 0 & 0 & \theta_{1} \\ \theta_{2} & 0 & 0 & 0 \\ 0 & \theta_{2} & 0 & 0 \\ 0 & 0 & \theta_{2} & 0 \\ 0 & 0 & 0 & \theta_{2}\end{array}\right], \quad$ and $\boldsymbol{\varphi}=\left[\begin{array}{l}\phi_{1} \\ \phi_{2} \\ \phi_{3} \\ \phi_{4}\end{array}\right]$
$\boldsymbol{\varepsilon}$ is a vector of random errors
The least squares system of normal equations is

$$
\begin{aligned}
& \left(\mathbf{x}^{T} \mathbf{x}\right) \hat{\boldsymbol{\varphi}}=\mathbf{x}^{T} \mathbf{y} \\
& \mathbf{x}^{T} \mathbf{x}=\left(\theta_{1}^{2}+\theta_{2}^{2}\right)\left[\begin{array}{llll}
1 & 0 & 0 & 0 \\
0 & 1 & 0 & 0 \\
0 & 0 & 1 & 0 \\
0 & 0 & 0 & 1
\end{array}\right] \\
& \mathbf{x}^{T} \mathbf{y}=\left[\begin{array}{l}
\theta_{1} y_{11}+\theta_{2} y_{21} \\
\theta_{1} y_{12}+\theta_{2} y_{22} \\
\theta_{1} y_{13}+\theta_{2} y_{23} \\
\theta_{1} y_{14}+\theta_{2} y_{24}
\end{array}\right]=\left[\begin{array}{c}
\sum_{i=1}^{2} \theta_{i} y_{i 1} \\
\sum_{i=1}^{2} \theta_{i} y_{i 2} \\
\sum_{i=1}^{2} \theta_{i} y_{i 3} \\
\sum_{i=1}^{2} \theta_{i} y_{i 4}
\end{array}\right]
\end{aligned}
$$

So $\hat{\boldsymbol{\varphi}}=\left(\mathbf{x}^{T} \mathbf{x}\right)^{-1} \mathbf{x}^{T} \mathbf{y}=\frac{1}{\sum_{i=1}^{2} \theta_{i}^{2}}\left[\begin{array}{l}\sum_{i=1}^{2} \theta_{i} y_{i 1} \\ \sum_{i=1}^{2} \theta_{i} y_{i 2} \\ \sum_{i=1}^{2} \theta_{i} y_{i 3} \\ \sum_{i=1}^{2} \theta_{i} y_{i 4}\end{array}\right]$
$\operatorname{Var}(\hat{\boldsymbol{\varphi}})=\sigma^{2}\left(\mathbf{x}^{T} \mathbf{x}\right)^{-1}=\frac{\sigma^{2}}{\sum_{i=1}^{2} \theta_{i}^{2}}\left[\begin{array}{llll}1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1\end{array}\right]$
Using step (a) to estimate $\boldsymbol{\theta}$ followed by step (b) to estimate $\boldsymbol{\varphi}$ iteratively are the basis of computing the least square estimates of the parameters using iterative fitting.

From $\hat{\boldsymbol{\theta}}=\left(\mathbf{x}^{T} \mathbf{x}\right)^{-1} \mathbf{x}^{T} \mathbf{y}=\frac{1}{\sum_{j=1}^{4} \phi_{j}^{2}}\left[\begin{array}{c}\sum_{j=1}^{4} \phi_{j} y_{1 j} \\ \sum_{j=1}^{4} \phi_{j} y_{2 j}\end{array}\right]$, making the constraint that the sum squares of $\phi s$ be equal to $J$ does not seem to make the model identifiable. It is seen that, in order to iteratively fit the set of $\theta$ s and $\phi s$, regarding the other set as known (as stated in the paper), we at least need starting values of the set of $\phi s$ (or the starting values of the set of $\theta s$ ).

Li and Wong stated that a large number ( $>10$ ) of arrays are needed for the probe sensitivity index $\phi s$ to be estimated accurately, otherwise the uncertainty in the estimation must be taken into account in the standard error computation.

